# A Markov chain identity and monotonicity of the diffusion constants for a random walk in a heterogeneous environment 

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(Received 5 October 1989; revised 9 November 1989)


#### Abstract

We consider a 2 -dimensional square lattice which is partitioned into a periodic array of rectangular cells, on which a nearest neighbour random walk with symmetric increments is defined whose transition probabilities only depend on the relative position within a cell. On the basis of a determinantal identity proved in this paper, we obtain a result for finite Markov chains which shows that the diffusion constants for the random walk are monotonic functions of the individual transition probabilities. We point out the similarity of this monotonicity property to Rayleigh's Monotonicity Law for electric networks or, equivalently, reversible random walks.


## 1. Introduction

Consider the following random walk on $\mathbb{Z}^{2}$ with site-dependent transition probabilities. Define a rectangular unit cell with $N$ sites to be the subset

$$
U=\left\{1,2, \ldots, m_{1}\right\} \times\left\{1,2, \ldots, m_{2}\right\} \subset \mathbb{Z}^{2}
$$

where $m_{1}$ and $m_{2}$ are integers denoting the width and height of the unit cell ( $m_{1} m_{2}=N$ ). A so-called 'inhomogeneously periodic' 2 -dimensional square lattice is obtained by repeating this unit cell periodically in two independent directions, such that the copies of the unit cell cover the grid $\mathbb{Z}^{2}$ without overlap (see Figure 1). The translation vectors $a_{1} \in \mathbb{Z}^{2}$ and $a_{2} \in \mathbb{Z}^{2}$ defining the partition into cells are not necessarily orthogonal; see Figure 2, where $\mathbf{a}_{1}=3 \mathbf{e}_{1}-\mathbf{e}_{2}, \mathbf{a}_{2}=3 \mathbf{e}_{2}$, with $\mathbf{e}_{1}$ and $\mathbf{e}_{2}$ the standard unit vectors in $\mathbb{Z}^{2}$. Any point $r$ on the lattice has a representation

$$
\mathbf{r}(l, \alpha)=l_{1} \mathbf{a}_{1}+l_{2} \mathbf{a}_{2}+\alpha_{1} \mathbf{e}_{1}+\alpha_{2} \mathbf{e}_{2}
$$

where the vector $l=\left(l_{1}, l_{2}\right) \in \mathbb{Z}^{2}$ serves to indicate the position of a translate of the unit cell and the vector $\alpha=\left(\alpha_{1}, \alpha_{2}\right) \in U$ denotes the relative position within a cell. To each site $\alpha$ in the unit cell assign a real variable $p_{\alpha}$, with $-\frac{1}{4}<p_{\alpha}<\frac{1}{4}$. A periodic environment with unit cell $U$ is the map $e: \mathbb{Z}^{2} \rightarrow\left\{p_{\alpha}: \alpha \in U\right\}$ given by

$$
\begin{array}{ll}
e(\mathbf{r})=e\left(\mathbf{r}+l_{1} \mathbf{a}_{\mathbf{1}}+l_{2} \mathbf{a}_{2}\right) & \left(\mathbf{r} \in \mathbb{Z}^{2}\right), \\
e(\mathbf{r})=p_{\mathbf{r}} & (\mathbf{r} \in U)
\end{array}
$$

Then a nearest neighbour random walk $X_{j} \in \mathbb{Z}^{2}$ with symmetric increments is defined by the transition probabilities

$$
P\left\{X_{j+1}-X_{j}= \pm \mathbf{e}_{1}\right\}=\frac{1}{2}-P\left\{X_{j+1}-X_{j}= \pm e_{2}\right\}=\frac{1}{4}+e\left(X_{j}\right)
$$

In other words, to each site $\alpha$ in the unit cell is assigned a transition probability $h_{\alpha}=\frac{1}{4}+p_{\alpha}$ in the positive and negative $x$-direction, and $\nu_{\alpha}=\frac{1}{4}-p_{\alpha}$ in the positive and negative $y$-direction (see Figure 3), and this assignment is repeated periodically. Sites with this type of transition probabilities have previously been termed anisotropic scatterers [6-9]. The site-symmetry in the positive and negative space directions causes the single-step averages of the horizontal and vertical displacements from every site to be zero. We call random walks with this property locally unbiased. Instead of infinite lattices we also consider finite lattices consisting of a single unit cell with periodic boundary conditions. In both cases we will speak of the 'anisotropic scatterer model'.

In our previous work on the anisotropic scatterer model, we have paid special attention to the diffusion constants, as defined below, for this model [7-9]. Explicit computation of these constants is only feasible for small size $N$ of the unit cell [7]. It is therefore of interest to establish qualitative properties of the diffusion constants, such as upper and lower bounds [9]. In this paper we will look at the dependence of the diffusion constants on the transition probabilities $h_{\alpha}$ and $\nu_{\alpha}$. It appears to be 'selfevident' that as the horizontal transition probability from one of the sites increases, diffusion in the horizontal direction increases as well. Our task will be to substantiate this intuition by mathematical proof.

We start by giving a definition of the diffusion constants. Both for the finite and the infinite lattice we define the total displacement $x(n)$ in the horizontal direction after $n$ steps by

$$
x(n)=\sum_{j=1}^{n} x_{j}
$$

where $x_{j}$ is 1 or -1 if the $j$ th step is in the positive or negative $x$-direction and zero otherwise. A similar definition holds for the total displacement $y(n)$ in the vertical direction. The horizontal and vertical diffusion constants of the random walk are defined by

$$
\begin{align*}
& D_{x}:=\frac{1}{2} \lim _{n \rightarrow \infty} \frac{1}{n}\left\langle x^{2}(n)\right\rangle \\
& D_{y}:=\frac{1}{2} \lim _{n \rightarrow \infty} \frac{1}{n}\left\langle y^{2}(n)\right\rangle
\end{align*}
$$

(see [6]), where the brackets denote an average over all realizations of the walk (notice that $\langle x(n)\rangle=\langle y(n)\rangle=0$, since the walk is locally unbiased).

The calculation of the diffusion constants proceeds as follows [6, 7]. For the finite lattice, let $T$ be the $N \times N$ transition matrix of the associated finite Markov chain with $N$ states. For the infinite inhomogeneously periodic lattice we define a similar matrix in the following way [6, 7]. Let the position of the walker be indexed by $(l, \alpha)$, where $l$ denotes the cell occupied by the walker and $\alpha$ the relative position within the cell ('internal state'). Let the single-step transition probability from site ( $l^{\prime}, \gamma$ ) to site $(l, \alpha)$ be denoted by $T_{\alpha \gamma}\left(l, l^{\prime}\right)$, which equals $T_{\alpha \gamma}\left(l-l^{\prime}\right)$ because of periodicity. Then an embedded (or lumped [2]) finite state Markov chain with $N$ states is constructed by


Fig. 1. Construction of a $2-\mathrm{D}$ lattice from a $3 \times 3$ unit cell with toroidal boundary

$$
\text { conditions: } a_{1}=(3,0), a_{2}=(0,3)
$$

ignoring the cell the walker occupies and only taking his internal state into account. The corresponding transition matrix $T$ has matrix elements $\dagger$

$$
T_{\alpha \gamma}=\sum_{l \in \mathbf{Z}^{2}} T_{\alpha \gamma}(l)
$$

In the sequel $T$ will always denote the transition matrix of the Markov chain on the finite lattice or, in the case of an infinite inhomogeneously periodic lattice, of the embedded Markov chain on the unit cell $U$. If the (embedded) Markov chain is ergodic there exists a unique stationary distribution vector $\pi$ which is the normalized right eigenvector $\dagger$ of the matrix $T$ corresponding to the eigenvalue $\lambda_{0}=1$. For the anisotropic scatterer problem as defined above ergodicity holds and the diffusion constants are given by,

$$
\begin{align*}
& D_{x}=\sum_{\alpha=1}^{N} h_{\alpha} \pi_{\alpha}=\frac{1}{4}+\sum_{\alpha=1}^{N} p_{\alpha} \pi_{\alpha} \\
& D_{y}=\sum_{\alpha=1}^{N} \nu_{\alpha} \pi_{\alpha}=\frac{1}{4}-\sum_{\alpha=1}^{N} p_{\alpha} \pi_{\alpha}
\end{align*}
$$

(see [7]).
To write down the transition matrix $T$ for the geometry of Figure 1, we label the sites in the unit cell by a double index ( $i j$ ), where $i$ runs in the horizontal and $j$ in the vertical direction ( $i=1,2, \ldots, m_{1}, j=1,2, \ldots, m_{2}$ ). Denoting the matrix entries by $\langle i j| T\left|i^{\prime} j^{\prime}\right\rangle$, we have

$$
\langle i j| T\left|i^{\prime} j^{\prime}\right\rangle=\left(\frac{1}{4}+p_{i^{\prime} j^{\prime}}\right)\left\{\delta_{i^{\prime}, i+1}+\delta_{i^{\prime}, i-1}\right\} \delta_{j, j^{\prime}}+\left(\frac{1}{4}-p_{i^{\prime} j}\right)\left\{\delta_{j^{\prime}, j+1}+\delta_{j^{\prime}, j-1}\right\} \delta_{i, i^{\prime}}
$$

(see [7]), where $i$ and $j$ are counted $\bmod m_{1}$ and $\bmod m_{2}$, respectively, because of the toroidal boundary conditions. For the case of Figure 2, where helical boundary

[^0]

Fig. 2. Construction of a 2-D lattice from a $3 \times 3$ unit cell with helical boundary conditions: $a_{1}=(3,-1), a_{2}=(0,3)$.
conditions are employed, it is easier to use a single index $i$ running from 1 to $N=$ $m_{1} m_{2}$. In this case the matrix entries are

$$
\langle i| T\left|i^{\prime}\right\rangle=\left(\frac{1}{4}+p_{i^{\prime}}\right)\left\{\delta_{i^{\prime}, i+1}+\delta_{i^{\prime}, i-1}\right\}+\left(\frac{1}{4}-p_{i^{\prime}}\right)\left\{\delta_{i^{\prime}, i+m_{1}}+\delta_{i^{\prime}, i-m_{1}}\right\},
$$

where now $i$ is counted $\bmod N$. Consider now the following
Question. If one of the horizontal transition probabilities is increased, does the horizontal diffusion constant $D_{x}$ increase too?

Notice that in the case of an inhomogeneously periodic lattice, increasing the transition probability at one site in the unit cell means increasing the transition probability at all the periodic translates of this site. On the basis of an explicit example we conjectured in [9] the validity of the following theorem, which we will prove in this paper:

Theorem 1•1. If any of the horizontal transition probabilities, say $h_{\alpha}=\frac{1}{4}+p_{\alpha}$, of the anisotropic scatterer model is increased, then the horizontal diffusion constant $D_{x}$ increases too. More precisely, for any $\alpha \in U$,

$$
\begin{equation*}
\frac{\partial D_{x}}{\partial h_{\alpha}}=N \pi_{\alpha}^{2} \tag{1.5}
\end{equation*}
$$

where $\pi$ is the stationary distribution of the corresponding finite state Markov chain.
Since the horizontal and vertical stepping probabilities are non-zero, ergodicity holds and the right-hand side of (1.5) is strictly positive, which is the monotonicity property alluded to above. From (1•2), $D_{x}+D_{y}=\frac{1}{2}$, so the corresponding results for the vertical diffusion constant are immediate and do not require separate discussion.

The relation ( 1.5 ) is similar to Rayleigh's Monotonicity Law for electric networks, which says that if the resistances of a circuit are increased, the effective resistance $R_{\text {EFF }}$ between any two points can only increase ; if they are decreased, it can only
decrease (see [1], p. 67). More precisely, let $R_{r s}$ be the resistance between sites $r$ and $s$ of the network, $I_{r s}$ the current flowing from $r$ to $s$ and $R_{\text {EFF }}$ the effective resistance between points $a$ and $b$ of the network. Under unit current flow, i.e. $I_{a}=-I_{b}=1$, where $I_{c}$ is defined as the sum of all currents leaving site $c$, it can be shown that

$$
\frac{\partial}{\partial R_{r s}} R_{\mathrm{EFF}}=I_{r s}^{2}
$$

(cf. [1], p. 77). We also note that Rayleigh's Monotonicity Law has an equivalent formulation in terms of escape probabilities for reversible random walks: see [1]. Both (1.5) and Rayleigh's Monotonicity Law establish a monotone dependence of macroscopic or 'effective' parameters (effective resistance, escape probability, diffusion constant) on the microscopic parameters of the model.

The purpose of this paper is to show that the Monotonicity Law (1.5) is a consequence of a Markov chain identity (Theorem $3 \cdot 2$ below), which in turn is based on a certain determinantal identity (Theorem $3 \cdot 1$ below) underlying all the results of this paper. This theorem holds for a general class of matrices including those of the form $S=I-T$, with $I$ the identity matrix and $T$ a stochastic matrix satisfying certain conditions. We will show in Section 3 that for the anisotropic scatterer problem with both types of boundary conditions used above, Theorem $3 \cdot 2$ is applicable and therefore ( 1.5 ) holds. Also certain types of waiting probabilities are allowed : one easily shows that the result ( 1.5 ) still holds if the transition probabilities are $a+b p_{\alpha}$ to jump horizontally from site $\alpha, c+d p_{\alpha}$ to jump vertically, and $1-2(a+c)-2(b+d) p_{\alpha}$ to stay at site $\alpha$, for $\alpha \in U$. Notice that all transitions from site $\alpha$ are described by a single parameter $p_{\alpha}$ (for an example with non-zero waiting probabilities where this condition is violated and ( 1.5 ) no longer holds, see [9]). It is straightforward to extend these results to higher dimensional lattices.

The monotonicity law ( 1.5 ) can be used to derive non-negativity of all derivatives of odd order of $D_{x}$ with respect to $p_{\alpha}$ : see [9]. It can also be applied to the case of random arrangements of the anisotropic scatterers inside the unit cell, where this random arrangement is then periodically translated as before. When $h_{\alpha}=a$ with probability $\rho$ and $h_{\alpha}=a^{\prime}$ with probability $1-\rho$, independently for all sites in the unit cell, it can be deduced from (1.5) that the average horizontal diffusion constant is an increasing function of the density $\rho$ when $a>a^{\prime}$. This result can be extended to the completely random lattice (i.e. infinite unit cell): see [9]. The existence of the diffusion constants for this case was proved by Lawler[3], who also obtained a rigorous low-density expansion [4].

The organization of the paper is as follows. Section 2 contains our notation and a summary of some prerequisites from the theory of determinants. In Section 3 we present a proof of the Markov chain identity mentioned above and we show that the result ( $1 \cdot 5$ ) is a special case of this. The basic determinantal identity (Theorem 3.1) is proved in Section 4.

## 2. Notation and elementary facts about determinants

In this section we establish our notation, mostly following Muir's treatise [5], and list some results about determinants. Matrices are denoted by capitals, $A, B, C$, etc.;


Fig. 3. Transition probabilities of the anisotropic scatterer model ; $\nu_{i}=\frac{1}{2}-h_{i}$.
matrix entries by lower case letters, e.g. $a_{i j}$ is the $(i, j)$-entry of the matrix $A$. The letter $I$ is reserved for the identity matrix. The columns of an $N \times N$ matrix $A$ are written in bold face: $\mathbf{a}_{1}, \mathbf{a}_{2}, \ldots, \mathbf{a}_{N}$. The transpose of a matrix $A$ is denoted by $A^{T}$, with a similar notation in the case of row or column vectors. The determinant of $A$ is denoted by $|A|$, so we will write

$$
|A|=\left|\begin{array}{cccc}
a_{11} & a_{12} & \cdots & a_{1 N} \\
a_{21} & a_{22} & \cdots & a_{2 N} \\
\cdot & \cdot & \cdots & \cdot \\
a_{N 1} & a_{N 2} & \cdots & a_{N N}
\end{array}\right|=\left|\mathbf{a}_{1}, \mathbf{a}_{2}, \ldots, \mathbf{a}_{N}\right|
$$

The determinant obtained from $|A|$ by deleting row $k$ and column $l$ is called a first minor, and denoted by $A_{k l}$. If $k=l$ we speak of a principal first minor. Minors in which the indices of the rows and columns taken to form the one are the same as the indices of the columns and rows taken to form the other are called conjugate minors. Multiplying the minor $A_{k l}$ by the sign factor $(-1)^{k+l}$ one obtains a first order (primary) cofactor of the element $a_{k l}$ in $|A|$, which is denoted by $\dagger$

$$
\mathscr{A}_{k l}=(-1)^{k+l} A_{k l}
$$

The following identity holds:

$$
\delta_{i j}|A|=\sum_{k} a_{i k} \mathscr{A}_{j k} .
$$

If we regard the determinant as a function of $N^{2}$ independent variables $\left\{a_{i j}\right\}$, we can express the cofactors as derivatives:

$$
\mathscr{A}_{i j}=\frac{\partial|A|}{\partial a_{i j}}
$$

Remark 2•1. In the sequel it will be tacitly understood that whenever derivatives of a determinant with respect to one or more of its entries occur, these entries are considered to be arbitrary (independent). Subsequently one may evaluate these derivatives for special values of the matrix entries, which then need no longer be independent.
$\dagger$ In [10] the cofactor $\mathscr{A}_{k l}$ is called a minor and denoted by $A_{l k}$.

In the following we denote by $e_{i}$ the $i$ th basis (column) vector of the standard basis in $\mathbb{R}^{N}$, defined by $\left(\mathbf{e}_{i}\right)_{j}=\delta_{i j}$. Also we write

$$
\mathrm{e}=(1,1, \ldots, 1)^{T}
$$

where the dimension of $e$ will be clear from the context. Then we have the following alternative notation for the cofactor $\mathscr{A}_{i j}$ :

$$
\mathscr{A}_{i j}=\left|\mathbf{a}_{1}, \ldots, \mathbf{a}_{j-1}, \mathbf{e}_{i}, \mathbf{a}_{j+1}, \ldots, \mathbf{a}_{N}\right|
$$

It will be tacitly understood that for $j=1$ the first column vector in expressions like $(2 \cdot 6)$ is $\mathbf{e}_{i}$. Both notations (2.4) and (2.6) will be used in the sequel. For example, the second order cofactor of the pair of elements $a_{i j}, a_{k l}$ in $|A|$, which is up to a sign factor equal to the determinant obtained from $|A|$ by deleting rows $i, k$ and columns $j, l$, can be written (assuming $j<l$ ) as

$$
\frac{\partial^{2}|A|}{\partial a_{i j} \partial a_{k l}}=\left|\mathbf{a}_{1}, \ldots, \mathbf{a}_{j-1}, \mathbf{e}_{i}, \mathbf{a}_{j+1}, \ldots, \mathbf{a}_{l-1}, \mathbf{e}_{k}, \mathbf{a}_{l+1}, \ldots, \mathbf{a}_{N}\right|
$$

Notice that this second derivative vanishes when $i=k$ or $j=l$.
Next we list a number of results which are needed below.
(i) Let each column-sum of $A$ be zero, i.e. $\Sigma_{i} a_{i j}=0$ for $j=1,2, \ldots, N$. Then

$$
|A|=0
$$

and

$$
\mathscr{A}_{j n}=\mathscr{A}_{n n} \quad \text { for all } j=1, \ldots, N
$$

i.e. the cofactors do not depend on the first index.

Proof. Adding rows $2, \ldots, N$ of $|A|$ to the first one, a row of zeros is obtained, hence $|A|=0$. To prove (2.7b) differentiate (2.3) with respect to $a_{m n}$ :

$$
\delta_{i j} \frac{\partial|A|}{\partial a_{m n}}=\delta_{i m} \mathscr{A}_{j n}+\sum_{k} a_{i k} \frac{\partial \mathscr{A}_{j k}}{\partial a_{m n}} .
$$

Now sum (2•8) over $i$ and use that $A$ has zero column-sums. Taking into account (2•4) one finds $\mathscr{A}_{m n}=\mathscr{A}_{j n}$, which proves $(2 \cdot 7 b)$. I
(ii) If the matrix $A$ is symmetric, i.e. $a_{i j}=a_{j i}$, conjugate minors are equal (see [5], p. 368), therefore

$$
\begin{equation*}
\mathscr{A}_{i j}=\mathscr{A}_{j i} . \tag{2.9}
\end{equation*}
$$

(iii) If $A$ both is symmetric and has zero column-sums then from (2.7) and (2.9),

$$
\mathscr{A}_{i j}=\mathscr{A}_{k l}, \quad \text { for all } i, j, k, l
$$

i.e. all primary cofactors are equal.
(iv) The following identity between the determinant of $A$ and its first and second order cofactors holds (see [5], p. 135):

$$
|A| \frac{\partial^{2}|A|}{\partial a_{i j} \partial a_{k l}}=\frac{\partial|A|}{\partial a_{i j}} \frac{\partial|A|}{\partial a_{k l}}-\frac{\partial|A|}{\partial a_{i l}} \frac{\partial|A|}{\partial a_{k j}} .
$$

(v) Differentiation of the relation (2-11) with respect to $a_{m n}$ gives

$$
\begin{align*}
& \frac{\partial \Delta}{\partial a_{m n}} \frac{\partial^{2} \Delta}{\partial a_{i j} \partial a_{k l}}+\Delta \frac{\partial^{3} \Delta}{\partial a_{m n} \partial a_{i j} \partial a_{k l}} \\
& \quad=\frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{i j}} \frac{\partial \Delta}{\partial a_{k l}}+\frac{\partial \Delta}{\partial a_{i j}} \frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{k l}}-\frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{i l}} \frac{\partial \Delta}{\partial a_{k j}}-\frac{\partial \Delta}{\partial a_{i l}} \frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{k j}},
\end{align*}
$$

where we have put $\Delta=|A|$. If $A$ has zero column-sums then from (2.7) this relation can be simplified to

$$
\frac{\partial \Delta}{\partial a_{n n}} \frac{\partial^{2} \Delta}{\partial a_{i j} \partial a_{k l}}-\frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{i j}} \frac{\partial \Delta}{\partial a_{l l}}=\frac{\partial \Delta}{\partial a_{j j}} \frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{k l}}-\frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{i l}} \frac{\partial \Delta}{\partial a_{j j}}-\frac{\partial \Delta}{\partial a_{l l}} \frac{\partial^{2} \Delta}{\partial a_{m n} \partial a_{k j}} .
$$

## 3. The main result

The transition matrices ( $1 \cdot 3$ ), ( $1 \cdot 4$ ) of the (embedded) Markov chain associated to the anisotropic scatterer model have the form $T=A^{\prime}+B^{\prime} P$. Here $A^{\prime}$ and $B^{\prime}$ are symmetric, (block) cyclic matrices and the column-sums of $A^{\prime}$ and $B^{\prime}$ are unity and zero, respectively, whereas $P$ is a diagonal matrix having the variables $\left\{p_{\alpha}\right\}$ appearing in Theorem $1 \cdot 1$ on its diagonal. Motivated by this, we prove in Section 4 the following theorem, from which all the results of this paper are immediate consequences. Notice that the theorem does not require the matrices involved to be cyclic.

Theorem 3•1. Let $S$ be an $N \times N$ matrix of the form $S=A(I+C P)$, where $I$ is the $N \times N$ identity matrix, $A$ and $C$ are symmetric $N \times N$ matrices with column-sums equal to zero, and $P$ is a diagonal matrix, say $(P)_{i j}=\delta_{i j} p_{j}$, where $p_{1}, \ldots, p_{N}$ are indeterminates. Let $S_{k k}, k=1,2, \ldots, N$, be the principal first minors of $|S|$, and define

Then

$$
Q=\sum_{k=1}^{N} S_{k k}, \quad R=\sum_{k=1}^{N} p_{k} S_{k k}
$$

$$
\begin{equation*}
\left(\frac{\partial R}{\partial p_{i}}\right) Q-R\left(\frac{\partial Q}{\partial p_{i}}\right)=N S_{i i}^{2} \quad(i=1,2, \ldots, N) \tag{3•2}
\end{equation*}
$$

Remark 3•1. The quantities $S_{k k}, Q, R$ occurring in this theorem are multilinear functions in the unspecified or formal variables $\left\{p_{i}\right\}$. The expression (3.2) therefore may be regarded as a succinct way of summarizing the set of identities between certain subdeterminants of $|S|$ which arise after expansion of (3•2) with respect to the variables $\left\{p_{i}\right\}$.

As an immediate corollary of Theorem $3 \cdot 1$ we have the following Markov chain identity.

Theorem 3.2. Let $T$ be the (column-stochastic) transition matrix of a Markov chain with $N$ states, and let $S:=I-T$, where $I$ is the $N \times N$ identity matrix. Assume that
(i) $S$ has the form $S=A(I+C P)$, where $A$ and $C$ are symmetric $N \times N$ matrices with column-sums equal to zero, and $P$ is a diagonal matrix, say $(P)_{i j}=\delta_{i j} p_{j}$, where $p_{1}, \ldots$, $p_{N}$ are real variables defined on a domain $\mathscr{X}$ in $\mathbb{R}^{N}$;
(ii) $\lambda_{0}=1$ is a simple root of $T$ for all values of $p_{1}, \ldots, p_{N}$ in $\mathscr{X}$.

Let $\pi$ be the stationary probability vector of the Markov chain, i.e. the right eigenvector of $T$ corresponding to the maximal root $\lambda_{0}=1$, and define $\mathscr{D}:=\sum_{k=1}^{N} p_{k} \pi_{k}$. Then for all values of $p_{1}, \ldots, p_{N}$ in $\mathscr{X}$,

$$
\frac{\partial \mathscr{D}}{\partial p_{i}}=N \pi_{i}^{2} \quad(i=1,2, \ldots, N) .
$$

Remark 3.2. Suppose that $S$ is as in Theorems $3 \cdot 1$ and $3 \cdot 2$, with the exception that the column-sums of $C$ are equal to some constant $K$, not necessarily zero. Then define a modified matrix $C^{*}$ by $c_{i j}^{*}=c_{i j}-K / N$. Clearly $C^{*}$ is symmetric, has zero columnsums and satisfies $A C^{*}=A C$, since

$$
\left(A C^{*}\right)_{i j}=\sum_{k} a_{i k} c_{k j}-\frac{K}{N} \sum_{k} a_{i k}=(A C)_{i j}
$$

where we have used the fact that the matrix $A$ has zero row-sums. So if $S$ has the form as in Theorems 3.1 and 3.2 where the matrix $C$ has constant column-sums, then the results still hold.

Remark 3.3. Condition (ii) is satisfied for ergodic chains, where the states form a single ergodic set (i.e. every state can be reached from every other state of the set: see [2], p. 37). In this case all components of $\pi$ are strictly positive. It is also satisfied when in addition to a single ergodic set there are transient states. Then the components $\pi_{\alpha}$ are positive for the ergodic and zero for the transient states.

Proof of Theorem 3.2. If $\lambda_{0}=1$ is a simple root of $T$, the components $\left\{\pi_{k}\right\}$ of the stationary vector $\pi$ can be expressed as,

$$
\pi_{k}=\frac{S_{k k}}{Q}
$$

where $S=I-T$. Here $S_{k k}$ for $k=1, \ldots, N$ are the principal first minors of $|S|$, which are non-negative and not all zero ([10], p. 21), and $Q$ is their sum. Therefore $\mathscr{D}$ can be written as $\mathscr{D}=R / Q$, where $R$ and $Q$ are defined in (3•1). Since

$$
\begin{equation*}
\frac{\partial \mathscr{D}}{\partial p_{i}}=\frac{\left(\partial R / \partial p_{i}\right) Q-R\left(\partial Q / \partial p_{i}\right)}{Q^{2}} \tag{3.5}
\end{equation*}
$$

and $S$ satisfies the conditions of Theorem 3.1, we find from (3.2),

$$
\frac{\partial \mathscr{D}}{\partial p_{i}}=N \frac{S_{i i}^{2}}{Q^{2}}=N \pi_{i}^{2} .
$$

Proof of Theorem 1•1. In the remainder of this section we will show that the anisotropic scatterer problem satisfies the conditions of Theorem 3.2 so that (3.6) and therefore ( 1.5 ) holds. Since the random walk is ergodic, condition (ii) of Theorem 3.2 is satisfied (see Remark 3.3) and we only have to demonstrate that the matrix $S=$ $I-T$ with $T$ defined by $(1 \cdot 3)$ or $(1 \cdot 4)$ can be written in the form required by condition (i) in Theorem 3.2.

First consider (1.4). We have $S=A+B P$, where the matrices $A$ and $B$ are given by

$$
\begin{gather*}
\langle i| A\left|i^{\prime}\right\rangle=\delta_{i, i^{\prime}}-\frac{1}{4}\left\{\delta_{i^{\prime}, i+1}+\delta_{i^{\prime}, i-1}\right\}-\frac{1}{4}\left\langle\delta_{i^{\prime}, i+m_{1}}+\delta_{i^{\prime}, i-m_{1}}\right\}, \\
\langle i| B\left|i^{\prime}\right\rangle=-\left\{\delta_{i^{\prime}, i+1}+\delta_{i^{\prime}, i-1}\right\}+\left\{\delta_{i^{\prime}, i+m_{1}}+\delta_{i^{\prime}, i-m_{1}}\right\} . \tag{3•8}
\end{gather*}
$$

It is clear that the matrices $A$ and $B$ are symmetric with zero column-sums. So we only have to show that $B=A C$ for some symmetric matrix $C$ with constant columnsums (see Remark 3•2).

Since the first minors of $A$ are non-zero (when all $p_{k}$ are zero, the walk is still ergodic), any set of $N-1$ columns of $A$ forms a basis for the space of $N$-dimensional vectors with zero column-sums. So in particular, any column $\mathbf{b}_{i}$ of $B$ can be uniquely written as a linear combination of the columns $\left\{\mathrm{a}_{k} ; k \neq i\right\}$ of $A$; say

$$
\begin{equation*}
\mathbf{b}_{i}=\sum_{k \neq i} c_{i k} \mathbf{a}_{k} \tag{3.9}
\end{equation*}
$$

We will show that the matrix $C$ occurring in (3.9) is symmetric with constant columnsums (the diagonal elements of $C$ are undetermined; for definiteness we set $c_{i i}=0$ ). First we observe that the matrices $A$ and $B$ are cyclic. Hence, if we denote by $W$ the permutation matrix which cyclically shifts every entry of a vector one position down, we can write

$$
\mathbf{a}_{k}=W^{k-1} \mathbf{a}_{1}, \quad \mathbf{b}_{k}=W^{k-1} \mathbf{b}_{1} \quad(k=1,2, \ldots, N)
$$

with $W^{N}=I$. From (3.9) there exist coefficients $\alpha_{k}$, for $k=2,3, \ldots, N$, such that

$$
\mathbf{b}_{1}=\sum_{k-2}^{N} \alpha_{k} \mathbf{a}_{k}
$$

hence by using (3.10) we find

$$
\mathbf{b}_{i}=\sum_{k=2}^{N} \alpha_{k}\left(W^{i-1} \mathbf{a}_{k}\right)=\sum_{k=2}^{N} \alpha_{k} \mathbf{a}_{k+i-1}=\sum_{k \neq i} \alpha_{k-i+1} \mathbf{a}_{k}
$$

where here and below indices are to be counted $\bmod N$. So the matrix $C$ has elements

$$
c_{i k}=\alpha_{k-i+1}
$$

where we put $\alpha_{1}:=0$. Clearly $C$ is a cyclic matrix, hence has constant column-sums. To show that $C$ is symmetric we invoke the site-symmetry of the anisotropic scatterer model in the horizontal and vertical directions, to which is associated the following symmetry operation. Let $U$ denote the matrix corresponding to the permutation $1 \rightarrow 1,2 \rightarrow N, 3 \rightarrow N-1, \ldots, N \rightarrow 2$. Under this permutation $\mathbf{a}_{1}, b_{1}$ are invariant, while $\mathbf{a}_{k} \rightarrow \mathbf{a}_{2-k}, \mathbf{b}_{k} \rightarrow \mathbf{b}_{2-k}$ (indices modulo $N$ ). Application of $U$ to (3.11) yields

$$
U \mathbf{b}_{1}=\sum_{k=2}^{N} \alpha_{k} U \mathbf{a}_{k}=\sum_{k=2}^{N} \alpha_{k} \mathbf{a}_{2-k}=\sum_{k=2}^{N} \alpha_{2-k} \mathbf{a}_{k}
$$

Since $U b_{1}=b_{1}$ and the vectors $\mathbf{a}_{2}, \ldots, a_{N}$ are independent, a comparison of (3.11) and (3.14) gives $\alpha_{k}=\alpha_{2-k}$, so that

$$
c_{i k}=\alpha_{k-i+1}=\alpha_{-k+i+1}=c_{k i}
$$

i.e. $C$ is symmetric. By Remark 3.2, Theorem $3 \cdot 2$ is applicable, so that from (1.2a),

$$
\frac{\partial D_{x}}{\partial h_{i}}=\frac{\partial \mathscr{P}}{\partial p_{i}}=N \pi_{i}^{2}
$$

where $\mathscr{D}$ is as in Theorem 3.2.

In the case of torus-geometry, Equation 1•3, one shows in the same way that the matrix $C$ is a symmetric cyclic block matrix, where the blocks are themselves symmetric and cyclic. Hence $C$ is again symmetric with constant column-sums. This completes the proof of the monotonicity law for the anisotropic scatterer problem.

## 4. Proof of Theorem 3•1

The proof of Theorem $3 \cdot 1$ is broken down into several steps, where the class of matrices involved is successively narrowed down. We start with the following lemma.

Lemma 4.1. Let $S$ be an $N \times N$ matrix of the form $S=A+B P$, where the matrices $A$ and $B$ have zero column-sums, and $P$ is a diagonal matrix, say $(P)_{i j}=\delta_{i j} p_{j}$ for $i, j=1$, $\ldots, N$. Let $\Xi^{(\alpha)}$ be defined as

$$
\Xi^{(\alpha)}=\left(\frac{\partial R}{\partial p_{\alpha}}\right) Q-R\left(\frac{\partial Q}{\partial p_{\alpha}}\right)
$$

where $Q$ and $R$ are given by (3.1) and $\alpha \in\{1,2, \ldots, N\}$. Then

$$
\Xi^{(\alpha)}=S_{\alpha \alpha}\left\{\sum_{k} S_{k k}-\sum_{k, l, i} p_{k} b_{i \alpha} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{i k}}\right\}
$$

Remark $4 \cdot 1$. The quantity $\Xi^{(\alpha)}$ is independent of $p_{\alpha}$. To see this notice that any first minor $S_{k k}$ is obviously independent of $p_{k}$ and (multi)linear in the other variables $p_{\beta}$ with $\beta \neq k$. So

$$
\begin{gathered}
R^{\prime}=S_{\alpha \alpha}+\sum_{k \neq \alpha} p_{k} S_{k k}^{\prime}, \quad R^{\prime \prime}=0 \\
Q^{\prime}=\sum_{k \neq \alpha} S_{k k}^{\prime}, \quad Q^{\prime \prime}=0
\end{gathered}
$$

where (double) primes denote (double) differentiation with respect to $p_{\alpha}$. Therefore $\Xi^{(\alpha)^{\prime}}=R^{\prime \prime} Q-R Q^{\prime \prime}=0$. Since $S_{\alpha \alpha}$ is itself independent of $p_{\alpha}$, the same is true for the quantity between braces in (4•2).

Proof of Lemma $4 \cdot 1$. From the definition of $Q$ and $R$ we have
with

$$
\begin{equation*}
\Xi^{(\alpha)}=S_{\alpha \alpha} \sum_{l} S_{l l}+\sum_{k, l} p_{k} \gamma_{k l} \tag{4•3}
\end{equation*}
$$

$$
\gamma_{k l}=S_{k k}^{\prime} S_{l l}-S_{k k} S_{l l}^{\prime}
$$

Since
we have

$$
\begin{gather*}
\frac{\partial}{\partial p_{\alpha}}=\sum_{i} b_{i \alpha} \frac{\partial}{\partial s_{i \alpha}}, \quad S_{k k}=\frac{\partial|S|}{\partial s_{k k}}, \\
\gamma_{k l}=\sum_{i} b_{i \alpha}\left\{\frac{\partial^{2}|S|}{\partial s_{i \alpha} \partial s_{k k}} \frac{\partial|S|}{\partial s_{l l}}-\frac{\partial^{2}|S|}{\partial s_{i \alpha} \partial s_{l l}} \frac{\partial|S|}{\partial s_{k k}}\right\} . \tag{4.5}
\end{gather*}
$$

Applying relation (2•13) to the matrix $S$, with $m=n \rightarrow l, j \rightarrow \alpha, l \rightarrow k$, we find

$$
\gamma_{k l}=\sum_{i} b_{i \alpha}\left\{S_{\alpha \alpha} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{k k}}-S_{\alpha \alpha} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{i k}}-S_{k k} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{k a}}\right\}=-\sum_{i} b_{i a} S_{\alpha a} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{i k}}
$$

where the second equality follows from the fact that $B$ has zero column-sums. A combination of (4.3) with (4.6) completes the proof.
To prove Theorem $3 \cdot 1$ we have to show that the quantity $\Xi^{(\alpha)}$ as defined by (4.1) is equal to $N S_{\alpha a}^{2}$. By applying the above lemma, which is clearly applicable in the setting of Theorem 3.1, it remains to be shown that $\Xi_{1}^{(\alpha)}=N S_{\alpha \alpha}$, where

$$
\Xi_{\mathrm{l}}^{(\alpha)}=\sum_{k} S_{k k}-\sum_{k, l, i} p_{k} b_{i z} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{k k}},
$$

is the quantity in braces in (4.2).
Our next step is to show that the dependence of $\Xi_{1}^{(\alpha)}$ on the matrix $A$ can be completely factored out. We first prove a simple lemma.
Lemma 42. Let $S$ be an $N \times N$ matrix of the form $S=A F$, where $A$ and $F$ are both $N \times N$ matrices and each row-sum of $A$ is zero. Then the following relation between the first order cofactors of the matrices $S, A$ and $F$ exists:

$$
S_{k k}=A_{k k} \sum_{m} \frac{\partial|F|}{\partial f_{m k}}
$$

Proof. In the notation introduced in (2.6) the right-hand side of (4.8) can be written as the product

$$
\left|\mathbf{a}_{1}, \ldots, \mathbf{a}_{k-1}, \mathbf{e}_{k}, \mathbf{a}_{k+1}, \ldots, \mathbf{a}_{N}\right|\left|\mathbf{f}_{1}, \ldots, \mathbf{f}_{k-1}, \mathbf{e}, \mathbf{f}_{k+1}, \ldots, \mathbf{f}_{N}\right|,
$$

where $\mathbf{e}$ is defined in (2:5). First we notice that, without affecting its value, we can replace the first factor in (4.9), which is the determinant of $A$ with the $k$ th column replaced by $\mathbf{e}_{k}$, by the determinant which arises by replacing the $k$ th row of $|A|$ by the transposed basis vector $\mathbf{e}_{k}^{T}$. Next we use the multiplication property of determinants to write (4.9) as a single determinant, say $|M|$, where $M$ is the product of the matrices involved in each of the factors. We find, using that $S=A F$ and $A$ has zero row-sums, that

$$
\begin{align*}
& m_{i j}=\sum_{l} a_{i t} f_{l j}=s_{i j} \quad(i \neq k, j \neq k) ; \\
& m_{k j}=f_{k j} \quad(j \neq k) ; \\
& m_{i k}=\sum_{l} a_{i l}=0 \quad(i \neq k) ; \\
& m_{k k}=1 .
\end{align*}
$$

We thus see that $|M|$ has the form

$$
|M|=\left|\mathbf{s}_{1}^{*}, \ldots, \mathbf{s}_{k-1}^{*}, \mathbf{e}_{k}, \mathbf{s}_{k+1}^{*}, \ldots, \mathbf{s}_{N}^{*}\right|,
$$

where $\mathbf{s}_{j}^{*}$ equals $\mathbf{s}_{j}$ except for the $k$ th entry. Since the $k$ th column of $M$ equals $\mathbf{e}_{k},|M|$ is clearly independent of the non-diagonal entries in the $k$ th row and therefore equals the minor $S_{k k}$ of $S$, which proves (4.8).
We now apply this lemma to the terms of $\Xi_{1}^{(\alpha)}$ in Equation 4.7, under additional conditions on the matrices $A$ and $B$ which are in accordance with the assumptions of Theorem 3•1.

Lemma 4.3. Let $S$ be an $N \times N$ matrix satisfying the conditions of Lemma $4 \cdot 1$, with the additional assumptions that $A$ is symmetric and $B=A C$ for some $N \times N$ matrix C. Define the matrix $D$ by $D:=I+C P$. Then the quantity $\Xi_{1}^{(\alpha)}$ in Equation 4.7 satisfies the identity

$$
\Xi_{1}^{(\alpha)}=\beta \Xi_{2}^{(\alpha)}
$$

where $\beta$ is the $k$-independent value of the first order cofactors $A_{k \boldsymbol{k}}$ of $A$, and

$$
\begin{gather*}
\Xi_{2}^{(\alpha)}=\sum_{k} D_{k}-\sum_{k, l, i, m} p_{k} c_{i \alpha} \frac{\partial^{2}|D|}{\partial d_{m l} \partial d_{i k}}, \\
D_{k}=\sum_{m} \frac{\partial|D|}{\partial d_{m k}}=\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{e}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{N}\right| .
\end{gather*}
$$

with

Proof. First, under the stated conditions, the matrix $S$ has the form $S=A D$, where $A$ is symmetric and has zero column-sums, therefore also zero row-sums. Applying Lemma $4 \cdot 2$ one finds

$$
S_{k k}=\beta D_{k}
$$

where $D_{k}$ is given by (4-13) and we have put $A_{k k}=\beta$, since the first order cofactors of $A$ do not depend on $k$; see ( $2 \cdot 10$ ). Next consider the second term in (4.7):

$$
\sum_{i} b_{i \alpha} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{i k}}=\left|\mathbf{s}_{1}, \ldots, \mathbf{s}_{k-1}, \mathbf{b}_{\alpha}, \mathbf{s}_{k+1}, \ldots, \mathbf{s}_{l-1}, \mathbf{e}_{l}, \mathbf{s}_{l+1}, \ldots, \mathbf{s}_{N}\right| .
$$

Since $S=A D$ and $B=A C$, this expression equals the minor $S_{l l}^{*}$ of the matrix $S^{*}=$ $A D^{*}$, where $D^{*}$ is obtained from $D$ by replacing the $k$ th column by $\mathrm{c}_{\alpha}$. So we can again apply Lemma $4 \cdot 2$ to find

$$
\sum_{i} b_{i \alpha} \frac{\partial^{2}|S|}{\partial s_{l l} \partial s_{i k}}=\beta\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{c}_{\alpha}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{l-1}, \mathbf{e}, \mathbf{d}_{l+1}, \ldots, \mathbf{d}_{N}\right|=\beta \sum_{i, m} c_{i \alpha} \frac{\partial^{2}|D|}{\partial d_{m l} \partial d_{i k}} .
$$

Substitution of $(4 \cdot 14)$ and $(4 \cdot 15)$ in (4.7) completes the proof. I
As our final step in the proof of Theorem $3 \cdot 1$ we have to show that $\Xi_{2}^{(\alpha)}$ equals $N D_{\alpha}$. This leads us to our last lemma.

Lemma 4.4. Let $C$ be a symmetric $N \times N$ matrix with zero column-sums, and let $D$ be the matrix $D=I+C P$, where $I$ is the unit matrix and $P$ is a diagonal matrix, say $(P)_{i j}=\delta_{i j} p_{j}$ for $i, j=1, \ldots, N$. Define

$$
\Xi_{2}^{(\alpha)}=\sum_{k} D_{k}-\sum_{k, l, i, m} p_{k} c_{i \alpha} \frac{\partial^{2}|D|}{\partial d_{m l} \partial d_{i k}}
$$

with $D_{k}$ given by (4-13). Then

$$
\Xi_{2}^{(\alpha)}=N D_{\alpha} .
$$

Remark $4 \cdot 2$. By Remark $4 \cdot 1$ the quantity $\Xi_{1}^{(\alpha)}$ in (4.7) or (4.11) is independent of $p_{\alpha}$. Since $\beta$ is a constant, (4.11) shows that the quantity $\Xi_{2}^{(\alpha)}$ in (4.16) is independent of $p_{\alpha}$ as well.

Remark 4.3. Without loss of generality we can take $\alpha=1$, for we can always
perform a permutation such that the $\alpha$ th row and column of all the matrices involved become the first row and column. Under such a permutation the right-hand side of $(4 \cdot 16)$ transforms to an expression of exactly the same form with $\alpha$ replaced by 1 and $C, P$ and $D$ replaced by matrices $C^{*}, P^{*}$ and $D^{*}$, having the same properties (symmetry etc.) as $C, P$ and $D$.

Proof of Lemma 4.4. We first show that

$$
\sum_{k} D_{k}=N|D|
$$

To see this, first notice that the sum in the left-hand side of (4-18) can be written as a single ( $N+1$ )-dimensional determinant

$$
\sum_{k} D_{k}=\sum_{k}\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{e}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{N}\right|=|\bar{D}|
$$

where

$$
|\bar{D}|=-\left|\binom{0}{\mathbf{e}},\binom{1}{\mathbf{d}_{\mathbf{1}}}, \ldots,\binom{1}{\mathbf{d}_{N}}\right| .
$$

This is easily checked by expanding $|\bar{D}|$ by its first row. Now subtract from the first row of $|\bar{D}|$ the sum of all the remaining rows, taking into account that under the assumptions of Lemma 4.4 the column-sum of any $d_{i}$ equals unity. The result is

$$
|\bar{D}|=-\left|\binom{-N}{\mathbf{e}},\binom{0}{\mathbf{d}_{1}}, \ldots,\binom{0}{\mathbf{d}_{N}}\right|=N\left|\mathbf{d}_{1}, \mathbf{d}_{2}, \ldots, \mathbf{d}_{N}\right|
$$

as was to be shown.
Applying the same reasoning to the second term in (4-16), using in addition that the column-sum of $c_{\alpha}$ is zero, one obtains

$$
\begin{align*}
\sum_{l, i, m} c_{i \alpha} \frac{\partial^{2}|D|}{\partial d_{m l} \partial d_{i k}} & =\sum_{\substack{l \\
l \neq k}}\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{c}_{\alpha}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{l-1}, \mathbf{e}, \mathbf{d}_{l+1}, \ldots, \mathbf{d}_{N}\right| \\
& =-\left|\binom{0}{\mathbf{e}},\binom{1}{\mathbf{d}_{1}}, \ldots,\binom{1}{\mathbf{d}_{k-1}},\binom{0}{\mathbf{c}_{\alpha}},\binom{1}{\mathbf{d}_{k+1}}, \ldots,\binom{1}{\mathbf{d}_{N}}\right| \\
& =N\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{c}_{\alpha}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{N}\right|
\end{align*}
$$

So we arrive at the result

$$
\Xi_{2}^{(\alpha)}=N \Xi_{3}^{(\alpha)}
$$

where

$$
\Xi_{3}^{(\alpha)}=\left|\mathbf{d}_{1}, \mathbf{d}_{2}, \ldots, \mathbf{d}_{N}\right|-\sum_{k=1}^{N} p_{k}\left|\mathbf{d}_{1}, \ldots, \mathbf{d}_{k-1}, \mathbf{c}_{\alpha}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{N}\right|
$$

The final step in the proof of Lemma $4 \cdot 4$ is to show that $\Xi_{3}^{(\alpha)}$ equals $D_{\alpha}$. By Remark $4 \cdot 2, \Xi_{3}^{(\alpha)}$ is independent of $p_{\alpha}$, so we are allowed to put $p_{\alpha}=0$ in (4.24), and by Remark $4 \cdot 3$ it is sufficient to give the proof for $\alpha=1$. Therefore consider

$$
\Xi_{3}^{(1)}=\left|\mathbf{e}_{1}, \mathbf{d}_{2}, \ldots, \mathbf{d}_{N}\right|+\mathscr{S}
$$

where

$$
\mathscr{S}=-\sum_{k=2}^{N} p_{k}\left|\mathbf{e}_{1}, \mathbf{d}_{2}, \ldots, \mathbf{d}_{k-1}, \mathbf{c}_{1}, \mathbf{d}_{k+1}, \ldots, \mathbf{d}_{N}\right|
$$

Now introduce the following notation for the column vector consisting of all except the first entry of $d_{i}$ :

$$
\mathrm{d}_{i}^{*}=\left(d_{2 i}, d_{3 i}, \ldots, d_{N i}\right)^{T}
$$

and similarly for $\mathrm{c}_{i}$, etc. Also write $C^{*}$ for the matrix obtained from $C$ by deleting the first row and column, and similarly $I^{*}$ for the corresponding truncation of the identity matrix $I$. Then one easily obtains

$$
\mathscr{S}=-\sum_{k=2}^{N} p_{k}\left|\mathbf{d}_{2}^{*}, \ldots, \mathbf{d}_{k-1}^{*}, \mathbf{c}_{1}^{*}, \mathbf{d}_{k+1}^{*}, \ldots, \mathbf{d}_{N}^{*}\right|=\left|\binom{0}{\mathbf{c}_{1}^{*}},\binom{p_{2}}{\mathbf{d}_{2}^{*}}, \ldots,\binom{p_{N}}{\mathbf{d}_{N}^{*}}\right|=\left|I_{1}+C_{1} P_{1}\right|,
$$

where $I_{1}, C_{1}$ and $P_{1}$ are the partitioned matrices

$$
I_{1}=\binom{0}{\hdashline I^{*}}, \quad C_{1}=\left(\begin{array}{c:c}
0 & \left(\mathrm{e}^{*}\right)^{T} \\
\hdashline \mathrm{c}_{1}^{*} & C^{*}
\end{array}\right), \quad P_{1}=\left(\begin{array}{c}
1 \\
\hdashline \\
\vdots P^{*}
\end{array}\right) .
$$

Now multiply in the last expression of (4-27) on the left by the determinant $\left|P_{1}\right|$ and on the right by $\left|P_{1}^{-1}\right|$ (since the variables $p_{2}, \ldots, p_{N}$ have only a formal meaning, we can assume without loss of generality that $P_{1}^{-1}$ exists). Then we obtain

$$
\mathscr{S}=\left|I_{1}+P_{1} C_{1}\right| .
$$

Next we use the invariance of a determinant under transposition to obtain

Finally we invoke the symmetry of the matrix $C$ :

$$
\left(C^{*}\right)^{T}=C^{\#}, \quad\left(\mathrm{c}_{1}^{*}\right)^{T}=\left(c_{21}, \ldots, c_{N 1}\right)=\left(c_{12}, \ldots, c_{1 N}\right)
$$

Substitution in (4.30) yields

$$
\mathscr{S}=\left|\binom{0}{\mathbf{e}^{*}},\binom{d_{12}}{\mathbf{d}_{2}^{*}}, \ldots,\binom{d_{1 N}}{\mathbf{d}_{N}^{*}}\right|=\left|\mathbf{e}-\mathbf{e}_{1}, \mathbf{d}_{2}, \ldots, \mathbf{d}_{N}\right| .
$$

To end the proof, insert (4-32) in (4.25) to find

$$
\Xi_{3}^{(1)}=\left|\mathbf{e}_{1}, \mathbf{d}_{2}, \ldots, d_{N}\right|+\left|e-e_{1}, d_{2}, \ldots, d_{N}\right|=\left|e, d_{2}, \ldots, d_{N}\right|=D_{1}
$$

So we conclude that $\Xi_{3}^{(\alpha)}=D_{\alpha}$ for all $\alpha$ and therefore, by $(4 \cdot 23), \Xi_{2}^{(\alpha)}=N D_{\alpha}$, as was to be shown.

Summarizing, by successive application of (4-2), (4.11) and (4.17) we find that

$$
\Xi^{(\alpha)}=S_{\alpha \alpha} \Xi_{1}^{(\alpha)}=S_{\alpha \alpha} \beta \Xi_{2}^{(\alpha)}=S_{\alpha \alpha} \beta N D_{\alpha}=N S_{\alpha \alpha}^{2},
$$

since $S_{\alpha \alpha}=\beta D_{\alpha}$ by (4-14). This proves the theorem.

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[^0]:    $\dagger$ We use column-stochastic matrices: $T_{\alpha \gamma}$ is the probability to go from state $\gamma$ to state $\alpha$, stationary distributions are right eigenvectors.

